

Class Notes for Chapter 8

- Survival data (i.e., ‘time until’ data) is typically **skewed** and **censored**. Censored data means right (most common), left, and/or interval. We’ll focus mostly on right-censored data analysis, although the nonlinear model in Section 8.6 addresses interval-censored data.
- Notation – need to get at an understanding of the survival and hazard functions. Survival function gives the probability of “survival” beyond some time-point. Hazard function gives the hazard/risk of ‘dying’ in the next instance given survival up to time t .
- The **Weibull distribution** is considered and illustrated here since it is useful in applications (especially reliability and engineering) and since it contains the **Exponential distribution** as a special case.
- Equation (8.7) is a key result – it gives the all-important LL!
- **Example 8.1 (Carcinoma)** – fit the Weibull distribution to these data; two values are right censored. MLE’s are given in Output 8.1a. Non-parametric (i.e., without any distributional assumption) estimate of the survival function: this estimate is called the **Kaplan-Meier estimate**. We can use SAS’ **LIFETEST** procedure to get it. Using both parametric and non-parametric methods, we obtain point and interval estimates of the median survival time (like LD_{50}) – see pp.5-6 for KM approach.
- **Example 8.2 (IUD)** – first graph is of N2LL for Exponential distribution, and second graph is obtained by profiling out the nuisance parameter (γ) and gives the N2PLL for the median survival time; intervals obtained by ‘cutting’.

- **Cox's Proportional Hazards (CPH) model** is used to relate hazard functions which vary with x , $h(t,x)$, to a baseline hazard (h_0) – the covariate vector (x) is assumed to enter as in Equation 8.8, or equivalently, as in Equation 8.9; covariates may include age, gender, etc. When there are two treatments (e.g., drugs A & B), x is just a dummy variable, and the slope parameter is given in Equation 8.10.
- Another model for survival data is the **Accelerated Failure Time [AFT] model** in Equation 8.12; it too brings covariate(s) into the model. Equations 8.12 and 8.13 are just for a dummy x , but can easily be extended to include other covariates.
- **Example 8.3 (Breast Cancer)** – the estimated survival curves on p.10 (obtained by KM estimation) appear to differ for the positively and negatively stained tumors: is this difference statistically significant or just an artifact?
 - (1) SAS **LIFETEST** procedure provides nonparametric estimates of quartiles and medians for the two groups, and, in Output 8.2c, three distinct tests of this question (null: no difference between the two curves). Many researchers use the **log-rank test**, but this may merely be by habit.
 - (2) SAS **PHREG** procedure fits the CPH model, gives an estimate of β in Output 8.2d, $\hat{\beta} = 0.9080$ – note that the likelihood test gives marginal significance here ($p = 0.0491$). The fitted result is given in Equation 8.11: the 'positive' survival curve is shifted down from the 'negative' survival curve. **Important interpretation: the hazard associated with the positive stain group is 2.479 times the hazard associated with the negative stain group – for all values of t .**

- (3) SAS **LIFEREG** (with **Weibull** distribution) fits the *parametric* Cox PH model – here the likelihood test that $\beta = 0$ gives $p = 0.0418$. *Note the doubling used in finding the TS here!*
- (4) SAS' **LIFEREG** (with **Log-logistic** distribution) fits the (parametric) AFT model – fitted results appear in Equation 8.14.
- Parametric, non-Normal SLR with censored data – use the LL given above to handle the censored measurements. For the Weibull distribution, median is more natural, so we use **median** in Equation 8.16 instead of **mean** in Equation 8.15.
 - [Example 8.4](#) – Y = time until death (no censored data here) and X is dose: wish to connect by a line. Intercept (β_0), slope (β_1) and variance-type parameter (γ) are estimated in Output 8.3a; the test of no significant relationship ($H_0: \beta_1 = 0$) is simple. Also simple is the test of LOF (lack of fit) of the assumed line since we could fit the full model: this is either the one-way ANOVA with 4 levels or a cubic polynomial. The ANOVA results are in Output 8.3b, and the line is fine here ($\chi^2_2 = 266.0 - 265.7 = 0.3$, NS).
 - **Interval-censored data** are addressed in [Example 8.5](#), so the binary logistic modelling ($\chi^2_1 = 31.1 - 26.4 = 4.7$, $p = 0.0302$) in Appendix 8.8.2 is wrong. Using the nonlinear model of Farrington (1996) and Collett (2003:286) in conjunction with the log-logistic distribution, we obtain Appendix 8.8.3 (p.22). Here the results are NS ($\chi^2_1 = 128.4 - 126.7 = 1.7$, $p = 0.1932$). We conclude there is no difference between drugs A and B.
 - Next issues:
 - nonlinear modelling with censored and skewed data
 - detecting synergy with censored and skewed data.