Maryland State Prison Data (Rossi et al (1980) and Allison (1995:4 & 42))

These data consist of a sample of n = 432 Maryland state prison inmates released from prison in 1980 and followed over a year until the first subsequent arrest.

Covariates include:

- age at release ('age' in years)
- employment status (dummy 'wexp' = 1 if the inmate had FT work experience before incarceration)
- financial assistance (dummy 'fin' = 1 if the inmate received financial aid after release and 0 otherwise); randomly assigned with equal numbers in each category
- marital status (dummy 'mar' = 1 if inmate was married at the time of release)
- parole status (dummy 'paro' = 1 if the inmate was released on parole)
- previous conviction status ('**prio**' = number of previous convictions)
- race (dummy 'race' = 1 if black and 0 otherwise)

In addition, the variable 'week' reflects the week of the first arrest (if applicable) and the variable 'arrest' is equal to 1 if 'week' is uncensored and equal to 0 if it is censored.

Fitting the Accelerated Failure Time (AFT) model with the Log-Logistic distribution, we obtain the following.

```
proc lifereg data=one;
  model week*arrest(0)=fin age race wexp mar
      paro prio / dist= dist=llogistic;
run;
```

	The LIFERE	G Procedure		
	Model In	formation		
	Data Set	WORK.	ONE	
	Dependent Variable	Log (we	ek)	
	Censoring Variable	arr	est	
	Censoring Value(s)		0	
	Number of Observati	ons	432	
	Noncensored Values		114	
	Right Censored Valu	es	318	
	Left Censored Value	S	0	
	Interval Censored V	alues	0	
	Name of Distributio	n LLogis	tic	
	Log Likelihood	-319.3983	3709	
	Number of Observati		32	
	Number of Observati	ons Used 4	32	
	Analysis of Par	ameter Estimates		
	Standard	95% Confidence	Chi-	
Parameter	DF Estimate Error	Limits	Square Pr > ChiSq	
Intercept	1 3.9183 0.4274	3.0805 4.7561	84.03 <.0001	
fin	1 0.2889 0.1456	0.0035 0.5742	3.94 0.0472	!
age	1 0.0364 0.0156	0.0058 0.0669	5.45 0.0195	•
race	1 -0.2791 0.2297	-0.7293 0.1710	1.48 0.2242	!
wexp	1 0.1784 0.1572	-0.1297 0.4865	1.29 0.2563	
mar	1 0.3473 0.2697	-0.1812 0.8758	1.66 0.1978	

paro	1	0.0508	0.1496	-0.2424	0.3440	0.12	0.7341
prio	1	-0.0692	0.0227	-0.1138	-0.0246	9.25	0.0023
Scale	1	0.6471	0.0559	0.5463	0.7666		

To interpret the parameter estimate for the 'fin' variable, since $\exp\{0.2889\} = 1.33$, after controlling for the other variables, the expected time to arrest for those who received financial assistance is 1.33 times the expected time to arrest for those who did not received financial assistance.

We should be careful with the above analysis as there may be (marginally) significant interaction between the financial assistance dummy variable and age; this means that the financial incentive program may impact younger and older inmates differently.

```
proc lifereg data=one;
  model week*arrest(0)=fin age race wexp mar
    paro prio fin*age / dist=llogistic;
run;
```

			Standard	95% Con	fidence	Chi-	
Parameter	DF	Estimate	Error	Lim		Square Pi	> ChiSq
Intercept	1	2.5159	1.4527	-0.3314	5.3631	3.00	0.0833
fin	1	-1.0222	0.6752	-2.3456	0.3012	2.29	0.1301
age	1	0.1655	0.1083	-0.0468	0.3777	2.33	0.1265
race	1	-0.2928	0.2298	-0.7432	0.1577	1.62	0.2027
wexp	1	0.1498	0.1586	-0.1610	0.4607	0.89	0.3447
mar	1	0.3034	0.2720	-0.2297	0.8366	1.24	0.2646
paro	1	0.0600	0.1498	-0.2336	0.3536	0.16	0.6887
prio	1	-0.0730	0.0227	-0.1176	-0.0284	10.31	0.0013
fin*age	1	0.0570	0.0290	0.0002	0.1138	3.88	0.0490
Scale	1	0.6435	0.0555	0.5433	0.7621		

Should we be interested in using the semi-parametric Cox Proportional Hazards model, we can use the following program and output.

```
The PHREG Procedure
                Model Information
        Data Set
                                 WORK.TWO
        Dependent Variable
                                 week
        Censoring Variable
                                 arrest
        Censoring Value(s)
        Ties Handling
                                 EFRON
     Number of Observations Read
                                          432
     Number of Observations Used
Summary of the Number of Event and Censored Values
                                         Percent
      Total
                           Censored
                                        Censored
                  Event
        432
                                318
                                           73.61
                    114
```

		Mo	odel Fit Stati	stics.		
			Without	: v	Vith	
		Criterion	Covariates	Covaria	ites	
		-2 LOG L	1350.761	1317.	495	
		AIC	1350.761	1331.	495	
		SBC	1350.761	1350.	649	
		Testing Glo	bal Null Hypo	thesis: BETA	\= 0	
	Tes [.]	t	Chi-Square	DF	Pr > ChiSq	
	Lik	elihood Ratio	33.2659	7	<.0001	
	Sco	re	33.5287	7	<.0001	
	Wal	t	32.1192	7	<.0001	
		Analysis of	F Maximum Like	lihood Estin	nates	
		Parameter	Standard			Hazard
Variable	DF	Estimate	Error	Chi-Square	Pr > ChiSq	Ratio
fin	1	-0.37942	0.19138	3.9304	0.0474	0.684
age	1	-0.05743	0.02200	6.8152	0.0090	0.944
race	1	0.31392	0.30799	1.0389	0.3081	1.369
wexp	1	-0.14981	0.21223	0.4983	0.4803	0.861
mar	1	-0.43372	0.38187	1.2900	0.2560	0.648
paro	1	-0.08486	0.19576	0.1879	0.6646	0.919
prio	4	0.09152	0.02865	10.2067	0.0014	1.096

For these data and this model, the estimate of the hazard ratio associated with the 'fin' variable is $\exp\{-0.37942\} = 0.684$; this means that the hazard of subsequent arrest for those who received financial assistance is 0.684 times the hazard of subsequent arrest for those who did not receive financial help. But again, we should be careful since 'fin*age' again appears to be significant (see below).

```
proc phreg data=two;
  model week*arrest(0)=fin age race wexp mar
  paro prio finage / ties=efron;
run;
```

		Т	he PHREG Proc	edure			
		Analysis	of Maximum Li	kelihood Estim	ates		
		Parameter	Standard			Hazard	
Variable	DF	Estimate	Error	Chi-Square	Pr > ChiSq	Ratio	
fin	1	1.62999	1.02908	2.5088	0.1132	5.104	
age	1	-0.02269	0.02596	0.7635	0.3822	0.978	
race	1	0.32129	0.30825	1.0864	0.2973	1.379	
wexp	1	-0.15865	0.21241	0.5579	0.4551	0.853	
mar	1	-0.45151	0.38272	1.3918	0.2381	0.637	
paro	1	-0.08384	0.19602	0.1829	0.6689	0.920	
prio	1	0.09421	0.02848	10.9418	0.0009	1.099	
finage	1	-0.08867	0.04508	3.8682	0.0492	0.915	

Quite different from the above survival analysis is the following <u>logistic analysis</u>. Here, the response variable is whether or not the inmate was subsequently arrested, and non-significant predictor variables have been removed from the model. Note that age enters this model in a quadratic manner.

```
proc logistic descending data=two;
  model arrest=fin age agesq prio educ fin*age;
run;
```

The LOGISTIC Procedure

Model Information

Data Set WORK.TWO
Response Variable arrest
Number of Response Levels 2

Model binary logit
Optimization Technique Fisher's scoring

Number of Observations Read 432 Number of Observations Used 432

Response Profile

Ordered			Total
Value	arrest		Frequency
1	1		114
2	0		318
Probability	modeled	is	arrest=1.

Model Fit Statistics

		Intercept
	Intercept	and
Criterion	Only	Covariates
AIC	500.602	475.234
SC	504.670	503.713
-2 Log L	498.602	461.234

Testing Global Null Hypothesis: BETA=0

Test	Chi-Square	DF	Pr > ChiSq
Likelihood Ratio	37.3681	6	<.0001
Score	34.7921	6	<.0001
Wald	30.7313	6	<.0001

Analysis of Maximum Likelihood Estimates

			Standard	Wald	
Parameter	DF	Estimate	Error	Chi-Square	Pr > ChiSq
Intercept	1	4.3304	2.2894	3.5779	0.0586
fin	1	1.6266	1.0674	2.3223	0.1275
age	1	-0.3223	0.1644	3.8429	0.0500
agesq	1	0.00513	0.00289	3.1598	0.0755
prio	1	0.0999	0.0371	7.2368	0.0071
educ	1	-0.2522	0.1538	2.6898	0.1010
fin*age	1	-0.0871	0.0452	3.7073	0.0542

Odds Ratio Estimates

	Point	95% Wald		
Effect	Estimate	Confidence	Limits	
agesq	1.005	0.999	1.011	
prio	1.105	1.027	1.188	
educ	0.777	0.575	1.050	